DS 100/200: Principles and Techniques of Data Science Date: March 8, 2019

Discussion #6 Exam Prep

Name:

1. Suppose in some universe, the true relationship between the measured luminosity of a single star Y can be written in terms of a single feature  $\phi$  of that same star as

$$Y = \theta^* \phi + \epsilon$$

where  $\phi \in \mathbb{R}$  is some non-random scalar feature,  $\theta^* \in \mathbb{R}$  is a non-random scalar parameter, and  $\epsilon$  is a random variable with  $\mathbb{E}[\epsilon] = 0$  and  $\operatorname{var}(\epsilon) = \sigma^2$ . For each star, you have a set of features  $\Phi = \begin{bmatrix} \phi_1 & \phi_2 & \dots & \phi_n \end{bmatrix}^T$  and luminosity measurements  $\mathbf{y} = \begin{bmatrix} y_1 & y_2 & \dots & y_n \end{bmatrix}^T$  generated by this relationship. Your  $\Phi$  may or may not include the feature  $\phi$  described above. The  $\epsilon_i$  for the various  $y_i$  have the same probability distribution and are independent of each other.

- (a) What is  $\mathbb{E}[Y]$ ?
  - $\bigcirc A. \ 0 \qquad \bigcirc B. \ \theta^* \phi \qquad \bigcirc C. \ \phi(\Phi^T \Phi)^{-1} \Phi^T \mathbf{y} \\ \bigcirc D. \ \theta^* \qquad \bigcirc E. \ \text{None of the above}$
- (b) What is var(Y)?

$$\bigcirc A. \quad \frac{\sigma^2}{n} \qquad \bigcirc B. \quad \frac{\sigma^2}{n^2} \qquad \bigcirc C. \quad 0$$
$$\bigcirc D. \quad \frac{1}{n-1} \sum_{i=1}^n \left( y_i - \frac{1}{n} \sum_{i=1}^n y_i \right)^2 \qquad \bigcirc E. \text{ None of the above}$$

2. What parameter estimate would minimize the following regularized loss function:

$$\ell(\theta) = \lambda(\theta - 4)^2 + \frac{1}{n} \sum_{i=1}^n (x_i - \theta)^2$$
(1)

$$\bigcirc A. \quad \theta = \frac{1}{\lambda n} \sum_{i=1}^{n} x_i$$
$$\bigcirc B. \quad \hat{\theta} = 4 + \frac{1}{\lambda n} \sum_{i=1}^{n} x_i$$
$$\bigcirc C. \quad \hat{\theta} = \frac{1}{n(\lambda+1)} \sum_{i=1}^{n} x_i$$
$$\bigcirc D. \quad \hat{\theta} = \frac{\lambda}{\lambda+1} + \frac{1}{n(\lambda+1)} \sum_{i=1}^{n} (x_i - 4)$$
$$\bigcirc E. \quad \hat{\theta} = \frac{4\lambda}{\lambda+1} + \frac{1}{n(\lambda+1)} \sum_{i=1}^{n} x_i$$

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3. Suppose  $X_1, \ldots, X_n$  are random variables with  $\mathbb{E}[X_i] = \mu^*$  and  $\operatorname{Var}[X_i] = \theta^*$ . Consider the following loss function

$$\ell(\theta) = \log(\theta) + \frac{1}{n\theta} \sum_{i=1}^{n} X_i^2.$$

Let  $\widehat{\theta}$  denote the minimizer for  $\ell(\theta)$ . What is  $\mathbb{E}[\widehat{\theta}]$ ?

 $\bigcirc A. \ \theta^* \ \bigcirc B. \ \theta^* + \mu^* \ \bigcirc C. \ \theta^* + \mu^*/2 \ \bigcirc D. \ \mathbb{E} \left[\theta^* + \mu^*\right] \ \bigcirc E. \ \theta^* + (\mu^*)^2$ 

- 4. Let  $x_1, \ldots, x_n$  denote any collection of numbers with average  $\overline{x} = \frac{1}{n} \sum_{i=1}^n x_i$ . (a)  $\sum_{i=1}^n (x_i \overline{x})^2 \leq \sum_{i=1}^n (x_i c)^2$  for all c.

 $\bigcirc$  A. True  $\bigcirc$  B. False

- (b)  $\sum_{i=1}^{n} |x_i \overline{x}| \le \sum_{i=1}^{n} |x_i c|$  for all *c*.  $\bigcirc$  A. True  $\bigcirc$  B. False
- 5. Consider the following loss function based on data  $x_1, \ldots, x_n$ :

$$\ell(\mu, \sigma) = \log(\sigma^2) + \frac{1}{n\sigma^2} \sum_{i=1}^n (x_i - \mu)^2.$$

- (a) Which estimator  $\hat{\mu}$  is a minimizer for  $\mu$ , i.e. satisfies  $\ell(\hat{\mu}, \sigma^2) \leq \ell(\mu, \sigma^2)$  for any  $\mu, \sigma?$ 
  - $\bigcirc$  A.  $\hat{\mu} = 0$  $\bigcirc B. \ \hat{\mu} = \frac{1}{n} \sum_{i=1}^{n} x_i$

$$\bigcirc C. \quad \hat{\mu} = \frac{1}{n} \sum_{i=1}^{n} x_i + \log\left(\frac{1}{n} \sum_{i=1}^{n} x_i\right)^2$$
$$\bigcirc D. \quad \hat{\mu} = \frac{1}{n\sigma^2} \sum_{i=1}^{n} x_i + \log\left(\sigma^2\right)$$
$$\bigcirc E. \quad \hat{\mu} = \texttt{median}(x_1, \dots, x_n).$$

(b) Which of the following is the result of solving  $\ell \sigma = 0$  for  $\sigma$  (for fixed  $\mu$ )?

$$\bigcirc A. \ \sigma = \frac{1}{n} \sum_{i=1}^{n} (x_i - \mu)^2.$$
  

$$\bigcirc B. \ \sigma = \sqrt{\frac{1}{n} \sum_{i=1}^{n} (x_i - \mu)^2}.$$
  

$$\bigcirc C. \ \sigma = \frac{2}{n} \sum_{i=1}^{n} (\mu - x_i).$$
  

$$\bigcirc D. \ \sigma = \sqrt{\frac{1}{n} \sum_{i=1}^{n} \sum_{j=1}^{n} (x_i - x_j)^2}.$$

6. Suppose we create a new loss function called the OINK loss, defined as follows for a single observation:

$$L_{OINK}(\theta, x, y) = \begin{cases} a(f_{\theta}(x) - y) & f_{\theta}(x) \ge y\\ b(y - f_{\theta}(x)) & f_{\theta}(x) < y \end{cases}$$

You decide to use the constant model (given on the left) and average OINK loss (given on the right).

$$f_{\theta}(x) = \theta$$
  $L(\theta, \mathbf{x}, \mathbf{y}) = \frac{1}{n} \sum_{i=1}^{n} L_{OINK}(\theta, x_i, y_i)$ 

The data are given below. Find the optimal  $\hat{\theta}$  that minimizes the loss.

- (a) when a = b = 1
- (b) when a = 1, b = 5
- (c) when a = 3, b = 6