DS 100/200: Principles and Techniques of Data Science Date: 10/25/19

Homework #6

Due Date: Friday, 11/1/19 at 11:59 PM

- You will turn in this homework by uploading your answers in PDF format to Gradescope. You may turn in your answer as a scan or good quality camera phone picture of handwritten sheets (e.g. CamScanner), or you may turn it in as a PDF generated from typeset math (e.g. using LaTeX or Microsoft Word).
- For this homework we have provided a companion notebook on DataHub. It is also linked on the course website. Problems 1, 8, and 9 explicitly ask you to run and interpret code in this notebook. You may also find the notebook helpful for problems 2 through 6. You will not need to turn in any .ipynb files.
- Due to resource constraints, we may elect to only grade a subset of the problems. We also may grade a subset of problems on completion, rather than on correctness. Since you dont know which problems these are, though, its in your best interest to fully attempt all of them.

## Collaborators

Data science is a collaborative activity. While you may talk with others about the homework, we ask that you write your solutions individually. If you do discuss the assignments with others please include their names at the top of your submission.

# Linear Regression Fundamentals and the Normal Equation

- 1. In this problem, we will review some of the core concepts in linear regression.
  - a. Suppose we create a linear model with parameters  $\hat{\beta} = [\hat{\beta}_0, \dots, \hat{\beta}_p]$ . As we saw in lecture, such a model makes predictions  $\hat{y} = \hat{\beta} \cdot \vec{x} = \sum \hat{\beta}_i x_i$ . Suppose  $\hat{\beta} = [2, 0, 1]$  and we receive an observation x = [1, 2, 3]. What  $\hat{y}$  value will this model predict for the given observation?
  - b. Suppose the correct y was 3.5. What will be the  $L_2$  loss for our prediction  $\hat{y}$  from question 1a?
  - c. In the companion notebook for this homework, we have provided a design matrix  $\mathbb{X}$  and a vector of response variables  $\hat{y}$ . These are given as variables  $\mathbf{X}$  and  $\mathbf{y}$  in the companion notebook. Suppose we create a linear regression model using this data. Explain briefly why  $\vec{\beta}$  will be a  $6 \times 1$  vector.
  - d. Using the normal equation from lecture 16, compute the optimal  $\hat{\beta}$ . Rather than giving all six values, in your answers, just tell us which of  $\hat{\beta}_0, \hat{\beta}_1, \hat{\beta}_2, \hat{\beta}_3, \hat{\beta}_4, \hat{\beta}_5$  is largest, and give its value rounded to two decimal places. Hint: np.linalg.inv might be useful. Hint: x.T gives the transpose of a matrix.
  - e. What will our model give for  $\hat{y}_1$ , i.e. what will it predict for the 1st observation (i.e. row one of X). Give your answer rounded to two decimal places.
  - f. What is the  $L_2$  loss for this prediction? What is the residual  $e_1$ ?
  - g. In lecture we said that the  $\hat{\beta}$  that results from the normal equation will "minimize the empirical risk". Does there exist a  $\vec{\beta}_{other}$  that would yield a lower loss for our prediction  $\hat{y}_1$ . If so, explain why we don't use  $\vec{\beta}_{other}$  instead. If not, explain why none exists.

Note: There are other equivalent notations for linear models. The notation we've used in this problem is consistent with what we saw in Prof. Nolan's lectures, i.e. using arrows to represent vectors, hats to represent estimates, and betas to represent parameters. Other sources use different notation, e.g. bolding for vectors or even not making any typographical distinction between vectors and scalars. Hats are often omitted. Some sources use the Greek letter  $\theta$  or the English letter w instead of  $\beta$ .

For instance, instead of saying  $\hat{y} = \vec{\beta} \cdot \vec{x}$ , the Data 100 textbook uses  $f_{\theta}(x) = \theta \cdot \mathbf{x}$ , i.e. bold to represent vectors, theta instead of beta, and  $f_{\theta}(x)$  instead of  $\hat{y}$ . CS 189 uses  $\hat{y}_i = w \cdot X_i$ . Data 100's Spring 2019 lectures used  $E[Y|X] = X^T \beta$ . Even our own lectures this semester are not entirely self consistent. We know it's annoying, but you'll just have to get used to this lack of a common consistent symbolic language.

#### **Observation Space vs. Variable Space**

2. The "variable space" approach views the design matrix X as a collection of  $n \times 1$  column vectors, one for each variable. On the other hand, the "observation space" approach considers the design matrix as a collection of  $1 \times p$  row vectors, one for each observation. In this exercise, we will examine many of the terms that we have been working with in regression (e.g.  $\hat{\beta}$ ) and connect them to their dimensions and to concepts that they represent. We will also draw connections between the observation and variable spaces.

First, we define some notation for the vectors in these two spaces. The  $n \times p$  design matrix  $\mathbb{X}$  corresponds to n observations on p variables (where one of these variables might actually be the one vector  $\vec{1}$ , a.k.a. a bias vector).  $\vec{y}$  is the response variable. We assume in this problem that we use  $\mathbb{X}$  and  $\vec{y}$  to compute optimal parameters  $\hat{\beta}$  for a linear model, and that this linear model generates predictions  $\vec{y}$  from  $\hat{\beta}$  and  $\mathbb{X}$  as we saw in lecture and in question 1 of this homework. We introduce new notation on this homework for the row and column vectors of  $\mathbb{X}$  as follows:

$$\vec{x}_{*j}$$
  $j^{th}$  column vector in  $\mathbb{X}, j = 1, \dots, p$   
 $\vec{x}_{i*}$   $i^{th}$  row vector in  $\mathbb{X}, i = 1, \dots, n$ 

Below, on the left, we have several expressions, labelled a through h, and on the right we have several terms, labelled 1 to 10. For each expression, determine its shape (e.g.,  $n \times p$ ), and match it to one the given terms. Terms may be used more than once or not at all. If a specific expression is nonsensical because the dimensions don't line up for a matrix multiplication, write "N/A" for both.

a. X 1. the residuals b.  $\vec{\hat{\beta}}$ 2. 0 3. 1st response,  $y_1$ c.  $\vec{x}_{*i}$ 4. 1st predicted value,  $\hat{y}_1$ d.  $\vec{x}_{1*}\hat{\beta}$ 5. 1st residual,  $e_1$ 6. the estimated coefficients e.  $\vec{x}_{*1}\hat{\beta}$ 7. the predicted values f.  $\mathbb{X}\hat{\hat{\beta}}$ 8. the features for a single observation 9. the value of a specific feature for all g.  $(\mathbb{X}^t \mathbb{X})^{-1} \mathbb{X}^t \vec{y}$ observations h.  $(I - \mathbb{X}(\mathbb{X}^t \mathbb{X})^{-1} \mathbb{X}^t) \vec{y}$ 10. the design matrix

As an example, for 2a, you would write: "2a. **Dimension:**  $n \times p$ , **Term:** 10".

#### Deriving Properties of the Simple Linear Regression

In lectures 14 and 15, we spent a great deal of time talking about the simple linear regression, which you also saw in Data 8. To briefly summarize, the simple linear regression model assumes that given a vector of observations  $\vec{x}$ , our predicted response for each of these observations is given by  $\vec{y} = \hat{\beta}_0 \vec{1} + \hat{\beta}_1 \vec{x}$ . Or equivalently, given a single observation x, our predicted response for this observation is  $\hat{y} = \hat{\beta}_0 + \hat{\beta}_1 x$ .

In lecture 14, we focused on the observation space representation, and saw that the  $\hat{\beta}_0$  and  $\hat{\beta}_1$  that minimize the  $L_2$  loss for the simple linear regression model are:

$$\hat{\beta}_0 = \bar{y} - \beta_1 \bar{x}$$
$$\hat{\beta}_1 = r \frac{SD_y}{SD_x}$$

Or, rearranging terms, our predictions  $\hat{y}$  are:

$$\hat{y} = \bar{y} + rSD_y \frac{x - \bar{x}}{SD_x}$$

In lecture 15, we used the exact same model, but now switched over to the variable space representation to get some geometric intuition for what we saw in lecture 14. The key geometric insight was that if we train a model on  $\vec{x}$  and  $\vec{y}$  and we use this model to make a prediction on a new observation x, our predicted  $\hat{y} = \hat{\beta}_0 x + \hat{\beta}_1$  is simply the vector in span $(\vec{1}, \vec{x})$  that is closest to y.

Consider some useful properties of the simple linear regression, listed below. Use the results derived from either the variable space or the observation space representation to prove these properties.

You may find the companion notebook helpful to support your thinking. See "Properties of Simple Linear Regression With and Without a Constant Term". Note: You may not answer questions 3 - 6 by simply computing the residuals of the given dataset and noting that the sum is zero. We want you to show that these properties are true for ALL possible datasets.

- 3. Show that/explain why the residuals from the fit have an average of 0, i.e.  $\sum e_i = 0$ .
- 4. Show that/explain why the  $n \times 1$  vectors  $\vec{x}$  and  $\vec{e}$  are orthogonal. In other words, explain why the dot product (a.k.a. inner product) of any observation used to train the model and the residuals is 0, i.e.  $\vec{x} \cdot \vec{e} = \sum (x_i e_i) = 0$ .
- 5. Show that/explain why the dot product of the residuals and  $\vec{y}$  is 0, i.e.  $\vec{y} \cdot \vec{e} = \sum (\hat{y}_i e_i) = 0$ .
- 6. Show that/explain why  $(\bar{x}, \bar{y})$  is on the regression line.

### Properties of a Linear Model With No Constant Term

Suppose that we don't include the intercept term in our model, that is, our model is now simply  $\vec{y} = \hat{\gamma}\vec{x}$ , where  $\hat{\gamma}$  is the single parameter for our model that we need to optimize. In this case, our least squares fit finds the  $\gamma$  that minimizes:

$$\sum_{i=1}^{n} (y_i - \gamma x_i)^2$$

for observed data  $(x_i, y_i), i = 1, \ldots, n$ .

The normal equations derived in lecture no longer hold. In this problem, we'll derive a solution to this simpler model. We'll see that the least squares estimate of the slope in this model differs from the simple linear regression model, and will also explore whether or not our properties from the previous problem still hold.

7. Use calculus to find the minimizing  $\hat{\gamma}$ . That is, prove that

$$\hat{\gamma} = \frac{\sum x_i y_i}{\sum x_i^2}$$

Note: This is the slope of our regression line, analogous to  $\hat{\beta}_1$  from our simple linear regression model.

- 8. In the previous section on deriving properties of estimators for a simple linear model, you established the following properties:
  - $\sum e_i = 0.$
  - $\vec{\hat{y}}$  and  $\vec{e}$  are orthogonal.
  - $\vec{x}$  and  $\vec{e}$  are orthogonal.
  - $(\bar{x}, \bar{y})$  is on the regression line.

Which of these properties are still true? Support your answers by giving the values of the following quantities as computed on the dataset given in "Properties of Simple Linear Regression With and Without a Constant Term" in the companion notebook for this homework.

- $\sum \bar{e}$
- $\vec{\hat{y}} \cdot \vec{e}$
- $\vec{x} \cdot \vec{e}$
- $\hat{\gamma}\bar{x}$

9. Recall that we can decompose the total sum of squares into two sum of squares, one measuring the variability "explained" by the regression and the other measuring the variability of the errors (the  $e_i$ ). Does this property still hold? That is,

$$\sum_{i} (y_i - \bar{y})^2 \stackrel{?}{=} \sum_{i} (y_i - \hat{y}_i)^2 + \sum_{i} (\hat{y}_i - \bar{y})^2$$

Support your answer by computing these three quantities using the dataset described in the previous problem.